**Supplement 1. Differences between responders and non-responders to open-ended questions.**

To assess whether there were any differences between participants who responded to at least one open-ended/qualitative question (64%, *n* = 300) and those who did not (36%, *n* = 169), we conducted *t*-tests using a variety of demographic and substantive variables as outcomes. There were no group differences for family income, child age, parent or child distress, child liking the United States, parental mediation, family dynamics, or participant race (i.e., White vs non-White). However, there were very small though statistically significant differences in education, parent age, reported COVID-19 impact and government trust. Responders had a higher education level than non-responders (1.92 vs. 2.14 on a scale of 1-7, where a lower number indicates a higher educational level, *p* < .05, Cohen’s *d* = .19) and were older (38.6 vs 37.6 years, *p* < .05, Cohen’s *d* = .23). Responders also reported higher COVID-19 impacts for both parent (3.88 vs. 3.59 on a scale of 1-5, *p <* .05, Cohen’s *d* = .28) and child (3.58 vs 3.21 on scale of 1-5, *p* < .001, Cohen’s *d* = .34), and lower trust in federal government (2.59 vs 3.10 on a scale of 1-7, *p* < .05, Cohen’s *d* = .31).

**Supplement 2. Content coding for mention of major current event.**

Altogether, 348 participants responded to one or both of the questions asking how other events besides the COVID-19 pandemic might be affecting them and/or their child substantively (i.e., anything other than “no” or an explanation of their child being too young). A dichotomous variable was created for whether participants did (=1) or did not (=0) mention the killing of Mr. George Floyd or subsequent community/national events in their response to either of the open-ended questions. The first author first coded each response into one of nine categories: eight categories included codes for any responses related to the killing of George Floyd, racism or inequality generally, or recent protests, riots, police brutality, or unrest all coded as 1. The final code was for “other,” which was coded as 0. To establish interrater reliability on the coding scheme, 20% of responses were randomly selected for a second researcher to code. The initial agreement across 864 possible codes (9 codes for 96 utterances) was 99%. After discussion of the discrepancies, the agreement was 100%.

**Supplement 3. Missing Data.**

The proportion of missing data was minimal across variables: 0.2% missingness across all PHQ-4 items; 0.2% missingness on White/non-White and federal government distrust; and 1.9% missingness for child disliking the U.S. Little’s MCAR Test was conducted to determine whether the data were missing completely at random (MCAR). The non-significant chi-square statistic indicates that the data is MCAR (χ2=112.93, *p*=0.54). Given the non-significant MCAR test and the very small amount of missing data in our sample, data were treated as MCAR and multivariate imputation by chained equations was used separately for the child and parent models. We conducted 100 imputations each for 10 participants missing one value across the variables in the child analysis and 3 participants missing one value across the variables in the parent analysis (‘mice’ package; Version 3.11.0; van Buuren & Groothuis-Oudshoorn, 2011). Analyses were also run using list-wise deletion and the pattern of results was similar. The descriptive statistics for Aim 1 are based on raw data, but multiple imputation was used for the Aim 2 analyses and statistics represent those using pooled datasets.

References

van Buuren, S., & Groothuis-Ousdhoorn, K. (2011). mice: Multivariate Imputation by Chained Equations in R. *Journal of Statistical Software,* *45*(3), 1-67. <https://www.jstatsoft.org/v45/i03/>